## Pion-Pion Scattering in the Non-Linear Sigma Model

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Ruaidhrí Campion Pion-Pion Scattering in the Non-Linear Sigma Model

#### Outline



- Ø Klein-Gordon Simulation
- On-Linear Sigma Model Simulation



Introduction & Background

Klein-Gordon Simulation Non-Linear Sigma Model Simulation Conclusion Lattice Field Theory Statistical Background

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Lattice Field Theory Statistical Background

## Klein-Gordon Theory

- Klein-Gordon theory<sup>1,2</sup> describes a non-interacting scalar boson
- Lattice action<sup>3</sup>

$$S(\phi) = \sum_{\mathbf{x}} \phi(\mathbf{x}) \left( \frac{2d + (am)^2}{2} \phi(\mathbf{x}) - \sum_{i=1}^d \phi(\mathbf{x} + \mathbf{e}_i) \right).$$

- d: lattice dimension
- a: lattice spacing
- m: boson mass

<sup>1</sup>Klein, Z. Phys. **37**, 895-906 (1926).
 <sup>2</sup>Gordon, Z. Phys. **40**, 117-133 (1926).
 <sup>3</sup>Rothe, Lattice Gauge Theories: An Introduction, 3rd ed., ch. 3 (2005).

Lattice Field Theory Statistical Background

## Non-Linear Sigma Model

- Non-linear sigma model<sup>4</sup> describes a scalar field  $\phi = (\sigma, \pi)$  taking values in a target manifold  $\mathcal{M} = S^{N-1}$
- N = 4: SU(2) NLSM/chiral model<sup>5</sup>
- Lattice action

$$S(\boldsymbol{\phi}) = -rac{eta}{2} \sum_{\mathbf{x}} \operatorname{Tr} \left[ \boldsymbol{\phi}(\mathbf{x}) \left( \lambda_0 \mathbb{I} + \sum_{i=1}^d \boldsymbol{\phi}^{\dagger}(\mathbf{x} + \mathbf{e}_i) \right) \right]$$

$$oldsymbol{\phi} = \sigma \mathbb{I} + i \sum_{k=1}^{3} \pi_k au_k \in SU(2)$$
  
det  $oldsymbol{\phi} = \sigma^2 + \pi^2 = 1 \implies (\sigma, \pi) \in S^4$ 

<sup>4</sup>Gellmann et al., *Nuovo Cim.* **16**, 705-726 (1960). <sup>5</sup>Gürsey, *Nuovo Cim.* **16**, 230-240 (1960). Introduction & Background

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## **Two-Point** Correlation

#### Define

$$egin{aligned} \Phi(t) &\equiv \sum_{i=1}^{d-1} \sum_{x_i=1}^{L_i} \phi((x_1,\ldots,\ldots,x_{d-1},t)) \ c(\delta) &\equiv \sum_{t=1}^T \Phi(t) \, \Phi(t+\delta) \end{aligned}$$

•  $c(\delta)$ : two-point correlation function<sup>6,7</sup>

General form<sup>8</sup>

$$c(\delta) \propto e^{-am\delta} + e^{-am(T-\delta)} + k$$

<sup>6</sup>Peskin et al., An Introduction to Quantum Field Theory, ch. 4.2 (1995).
<sup>7</sup>Rothe, Lattice Gauge Theories: An Introduction, 3rd ed., ch. 2 (2005).
<sup>8</sup>Rothe, Lattice Gauge Theories: An Introduction, 3rd ed., ch. 10.1 (2005).

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Lattice Field Theory Statistical Background

## Markov Chain Monte Carlo Methods

- Randomly generate a sequence of states independent of previous states to calculate expected values for some distribution  $\lambda(\phi)$
- Metropolis-Hastings algorithm<sup>9,10,11</sup>
  - Current state  $\nu$
  - Propose state  $\xi$  according to some  $q(\xi, \nu)$
  - Accept  $\xi$  as next state with probability

$$\alpha(\xi,\nu) = \min\left(\frac{q(\nu,\xi)\lambda(\xi)}{q(\xi,\nu)\lambda(\nu)},1\right)$$

• Gibbs sampler:<sup>12,11</sup>  $q(\xi, \nu) \propto \lambda(\xi) \implies \alpha = 100\%$ 

<sup>9</sup>Metropolis et al., *J. Chem. Phys.* **21**, 1087-1092 (1953). <sup>10</sup>Hastings, *Biometrika* **57**, 97-109 (1970). <sup>11</sup>Ross, *Simulation*, 5th ed., ch. 12 (2005). <sup>12</sup>Geman et al., *IEEE Trans. Pattern Anal. Mach. Intell.* **6**, 721-741 (1984).

Lattice Field Theory Statistical Background

#### Autocorrelation

- Initial state far from equilibrium: discard first D states
- Naïve  $MSE = \frac{variance}{no. of samples} \neq true MSE$ 
  - Binning method:<sup>13</sup> average samples into bins, binned MSE =  $\frac{\text{variance of binned samples}}{\text{no. of binned samples}} \rightarrow \text{true MSE}$  as bin size  $\rightarrow \infty$
  - Integrated autocorrelation time  $^{14}$  for estimator  $\hat{\mu}$

 $\tau_{\hat{\mu}} = \frac{\text{true MSE}}{\text{naïve MSE}} = \frac{\text{total no. of samples}}{\text{effective no. of ind. samples}}$ 

<sup>13</sup>Knechtli et al., *Lattice Quantum Chromodynamics...*, ch. 2 (2017).
 <sup>14</sup>Sokal, Cargése Summer School (1996).

Lattice Field Theory Statistical Background

## **Error** Calculation

Random variables  $A_1, \ldots, A_N$  with mean  $\mu$ 

- Sample mean  $\hat{\mu}$ : binning method
- Estimating f(µ)
  - Average of  $f(A_i)$  does not converge to  $f(\mu)^{15}$
  - Jackknife method<sup>16,17</sup> for independent samples

$$\hat{f}_{i} = f\left(\frac{1}{N-1}\sum_{k\neq i}A_{k}\right),$$
$$\hat{f}_{jack} = \frac{1}{N}\sum_{i=1}^{N}\hat{f}_{i}, \quad \widehat{\mathsf{MSE}}(\hat{f}) = \frac{N-1}{N}\sum_{i=1}^{N}\left(\hat{f}_{i} - \hat{f}_{jack}\right)^{2}$$

Dependent samples: combine jackknife & binning
 <sup>15</sup>Berg, Markov Chain Monte Carlo Simulations..., ch. 2.7 (2004).
 <sup>16</sup>Quenouille, Biometrika 43, 353-360 (1956).
 <sup>17</sup>Tukey, Ann. Math. Stat. 29, 614 (1958).

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Motivation

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- Debug & verify calculations by analytic comparison
- Optimising approaches to writing programmes
- Compare efficiency of different MCMC algorithms
- Determine relationships between  $\tau_{\hat{c}(\delta)}$  and system parameters (lattice size, dimensionless mass *am*)

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#### Analytic Comparison



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#### Binning Method



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#### Metropolis-Hastings Algorithm Procedure

• Calculate 
$$\kappa^2 = \frac{1}{2d + (am)^2}$$
.

2 Choose a lattice site **x** and calculate 
$$\gamma(\mathbf{x}) = \sum_{i=1}^{d} (\phi(\mathbf{x} - \mathbf{e}_i) + \phi(\mathbf{x} + \mathbf{e}_i)).$$

- 3 Generate a proposed value  $\phi'(\mathbf{x}) \sim \phi(\mathbf{x}) + \mathcal{U}(-\varepsilon, \varepsilon)$ .
- Calculate  $\psi(\mathbf{x}) = \phi(\mathbf{x}) \kappa^2 \gamma(\mathbf{x})$  and  $\psi'(\mathbf{x}) = \phi'(\mathbf{x}) \kappa^2 \gamma(\mathbf{x})$ .
- If |ψ'(x)| ≤ |ψ(x)|, then set φ(x) = φ'(x). Otherwise, generate U ~ U(0, 1) and set φ(x) = φ'(x) if U < exp [-(ψ'(x))<sup>2</sup>-(ψ(x))<sup>2</sup>].
- Q Repeat steps 2-5 for each lattice site.

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## Gibbs Sampler Procedure

• Calculate 
$$\kappa^2 = \frac{1}{2d + (am)^2}$$
.

- **2** Choose a lattice site **x** and calculate  $\gamma(\mathbf{x})$ .
- **③** Generate a proposed value  $\phi'(\mathbf{x}) \sim \mathcal{N}(\kappa^2 \gamma(\mathbf{x}), \kappa^2)$ .

• Set 
$$\phi(\mathbf{x}) = \phi'(\mathbf{x})$$
.

Q Repeat steps 2-4 for each lattice site.

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#### Autocorrelation



• Metropolis-Hastings: optimal  $\hat{\tau}_{\hat{c}(\delta)} = 63.21$  for  $\varepsilon = 1.58$ 

• Gibbs: 
$$\hat{ au}_{\hat{c}(\delta)} = 17.75$$

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## Lattice Size



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#### Mass



•  $au_{\hat{c}(\delta)} - 1 \propto am^{-2}$  for am < 1

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## Procedure

Or Calculate  $v = v_0 + i \sum_{k=1}^{3} v_k \tau_k$  and set  $\phi(\mathbf{x}) = v \cdot \frac{\mathbf{\Sigma}(\mathbf{x})}{\sqrt{\det \mathbf{\Sigma}(\mathbf{x})}}$ .

• Repeat steps 1-5 for each lattice site.

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## **Debugging & Verification**



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## Debugging & Verification



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## Debugging & Verification



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Gibbs Sampler **Observable Calculation** Masses Autocorrelations

#### Two-Point Correlations





•  $c_{\pi}(\delta) = \sum_{i=1}^{3} c_{\pi_i}(\delta)$ 

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#### Mass

• 
$$c(\delta) = e^{-am\delta} + e^{-am(T-\delta)} + k$$
  
• Small  $\delta \implies c(\delta) \approx k_1 e^{-am\delta} + k_2$   
• ...  
•  $am \approx \ln\left\{\frac{c(0)-c(2)+\sqrt{[c(0)-c(2)]^2-4[c(1)-c(2)][c(0)-c(1)]}}{2[c(1)-c(2)]}\right\}$   
•  $am(c(\delta)) \implies \widehat{am}$  using jackknife

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- Masses independent of lattice size
- Autocorrelation times maximised when mass minimised
- $am_{\sigma}$  has a  $\lambda_0$ -dependent minimum

• 
$$(am_\pi)^2 o \lambda_0$$
 as  $eta o \infty$ 

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# $au_{\widehat{am}}$ and $au_{\widehat{c}(\delta)}$



- $\tau_{\widehat{am}} < \tau_{\widehat{c}(\delta)}$
- $\tau_{\sigma} < \tau_{\pi}$
- ... Algorithm efficiency depends on observable

Project Summary

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Project Summary

## **Project Summary**

- Klein-Gordon theory
  - Verified results via analytic comparison & statistical theory
  - Found Gibbs > Metropolis-Hastings
  - Autocorrelation depends on am, independent of lattice size
- Non-linear sigma model
  - Used code and results from Klein-Gordon study
  - Investigated particle mass behaviour on system parameters
  - Observable determines algorithm efficiency

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