

Fourier transform of the probability measures

Romeo Vomigescu

Abstract

In this paper we make the connection between Fourier transform of a probability measure and the characteristic function in the \mathbb{R}^2 space; also we establish some the properties.

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1. Let Ω and Ω' be any sets, let \mathcal{K} and \mathcal{K}' to be two σ -algebras on Ω and Ω' respectively and the measurable spaces (Ω, \mathcal{K}) , (Ω', \mathcal{K}') .

A function $f : \Omega \rightarrow \Omega'$ is said to be $(\mathcal{K}, \mathcal{K}')$ - measurable if the borelian filed $f^{-1}(\mathcal{K}') \subset \mathcal{K}$. Let $f : (\Omega, \mathcal{K}) \rightarrow (\Omega', \mathcal{K}')$ be a measurable function and let $\mu : \mathcal{K} \rightarrow [0, \infty)$ be a measure on \mathcal{K} . Then the function of set $\mu \circ f^{-1}$ defined on \mathcal{K}' by the rule $\mu \circ f^{-1}$ is called, the image of the measure μ by f .

The triple $(\Omega, \mathcal{K}, \mu)$ where (Ω, \mathcal{K}) is a measurable space, and μ is a measure on \mathcal{K} , is called the space with measure. If $\mu(\Omega) = 1$, then μ is called the probability measure.

Let $\varphi : (\Omega, \mathcal{K}) \rightarrow (\Omega', \mathcal{K}')$ be a measurable function $f : (\Omega, \mathcal{K}) \rightarrow X(\mathbb{R} \vee \mathbb{C})$ is $\mu \circ \varphi^{-1}$ - integrable, if and only if $f \circ \varphi$ is μ -integrable. In this case the following relation holds

$$(1) \quad \int (f \circ \varphi) d\mu = \int f \cdot d\mu \circ \varphi^{-1}$$

and is called the transport formula.

2. We note with (Ω, \mathcal{K}, P) a probability field and let (X, \mathcal{X}) be a measurable space where \mathcal{X} is a borelian field on X . A measurable functions $f : (\Omega, \mathcal{K}, P) \rightarrow (X, \mathcal{X})$ is called a random variable. If the function f is a random variable, then the image of P by f we will note with $P \circ f^{-1}$ and will be called, distribution of f . In this case the distribution of f is the probability on \mathcal{X} defined by $P \circ f^{-1}(A) = P(f^{-1}(A))$, $A \in \mathcal{X}$. This events $f^{-1}(A)$ we also denote by $\{f \in A\}$. If F is a probability on \mathbb{R}^k , then we will say that F has the density ρ if $F < m_k$ (m_k is the Lebesque measure on \mathbb{R}^k) and ρ is a version of the Radonikodym derivative dF/dm .

(For λ, μ -measures, $\lambda < \mu$ denote that λ is absolutely continuous with respect to μ , i.e. $\mu(A) = 0$ implies that $\lambda(A) = 0$)

If the function $f : (\Omega, \mathcal{K}, P) \rightarrow \mathbb{R}^k$ is a random variable, we will say that f has the density ρ if the distribution $P \circ f^{-1}$ has the density ρ . Hence a function $\rho : \mathbb{R}^k \rightarrow \overline{\mathbb{R}}$ is the density of the random variable f if:

- i) ρ is measurable and $\rho \geq 0$
- ii) $P(f \in A) = \int_A \rho(x) dm_k(x)$, $A \in \mathcal{B}_{\mathbb{R}^k}$, where \mathcal{B} is a borelian field.

For a random variable $f : (\Omega, \mathcal{K}, P) \rightarrow \mathbb{R}^k$ and for a measurable function $\varphi : \mathbb{R}^k$ and for a measurable function $\varphi : \mathbb{R}^k \rightarrow \mathbb{C}$, the transport formula can be written as

$$(2) \quad \int_{\Omega} \varphi \circ f dP = \int_{\mathbb{R}^k} \varphi(x) dP \circ f^{-1}(x)$$

In particular, if f has the density ρ , then

$$(3) \quad \int_{\Omega} \varphi \circ f dP = \int_{\mathbb{R}^k} \varphi(x) \rho(x) dx$$

Let $\zeta = (\xi, \eta)$ be a random vector whose components are the random variables ξ and η . If so, the function F define by the relation

$$(4) \quad F(z) = F(x, y) = P(\xi \leq x, \eta \leq y), \quad \forall(z) = (x, y) \in \mathbb{R}^2$$

is called the distribution function of the random vector ζ , where $P(\xi \leq x, \eta \leq y)$ is the probability that an aleatory point $\xi \in (-\infty, x]$, $\eta \in (-\infty, y]$.

The function F has analogous properties with the distribution function from the unidimensional case:

$$0 \leq F(x, y) \leq 1, \quad \lim_{x, y \rightarrow -\infty} F(x, y) = 0, \quad \lim_{x, y \rightarrow \infty} F(x, y) = 1.$$

The monotony condition of the function F will be characterized by the following inequalities:

$$F(x + h, y) - F(x, y) \geq 0, \quad F(x, y + h) - F(x, y) \geq 0$$

$$F(x + h, y + h) - F(x + h, y) \geq F(x, y + h) - F(x, y)$$

where h and k represent two positive increases. Let \mathcal{V} be boolean algebra of all B-intervals of the form

$$\Delta = [a, b] \times [c, d], \quad a, b, c, d \in \mathbb{R}$$

and let $\mu : \mathcal{V} \rightarrow [0, \infty]$ be a measure on \mathcal{V} so that $\mu(\Delta) < \infty$.

We know (see [3]) that there exists a monotone nondecreasing and left-continuous function F on \mathbb{R}^2 , so that $\forall a, b, c, d \in \mathbb{R}$ we have

$$(5) \quad \mu([a, b] \times [c, d]) = F(b, d) - F(a, d) - F(b, c) + F(a, c) = P(\zeta \in \Delta)$$

The reciprocal being also valid.

If F_1 and F_2 are monotone non-decreasing and left-continuous functions on \mathbb{R}^2 , so that

$$\begin{aligned} \mu([a, b] \times [c, d]) &= F_1(b, d) - F_1(a, d) - F_1(b, c) + F_1(a, c) = \\ &= F_2(b, d) - F_2(a, d) - F_2(b, c) + F_2(a, c), \quad \forall a, b, c, d \in \mathbb{R}, \end{aligned}$$

then there exists a hyperbolic constant.

$\psi(x, y) = \varphi(x) + \psi(y)$ so that $F_2(x, y) = F_1(x, y) + \psi$. If μ is a measure on \mathcal{V} with $\mu(\mathbb{R}^2) = \alpha < \infty$, then a monotone non-decreasing and left-continuous function F on \mathbb{R}^2 , can be found, having the properties

$$\lim_{x, y \rightarrow -\infty} F(x, y) = 0, \quad \lim_{x, y \rightarrow \infty} F(x, y) = \alpha \text{ and (4) holds.}$$

The function F so defined, is unique. If $\alpha = 1$, then the function F is called distribution (probability).

3. Let $\zeta = (\xi, \eta)$ be a random vector. Then, one defines for each measure μ on \mathbb{R}^2 , Fourier transform or otherwise characteristicly function of the probability measure

$$(6) \quad \widehat{\mu}(t) = \int e^{i\langle t, z \rangle} dF(z), \quad t \in \mathbb{R}^2$$

where $t = (u, v)$, $z = (x, y) \in \mathbb{R}^2$. This function is called the distribution of μ . We have

$$(7) \quad \widehat{\mu}(t) = \int_{\mathbb{R}^2} e^{i\langle t, z \rangle} dF(z)$$

where $F(z)$ has the expression (4).

If the random variable $f : (\Omega, \mathcal{K}, P) \rightarrow \mathbb{R}^2$ and $\mu = P \circ f^{-1}$ is distribution of f , then the characteristicly function $\widehat{\mu}$ is

$$(8) \quad \widehat{\mu}(t) = \int_{\Omega} e^{i\langle t, z \rangle} d\mu(z) = \int_{\Omega} e^{i\langle t, f \rangle} dP = M \cdot e^{i\langle t, f \rangle}$$

where M is the mean value. In this case we say that $\widehat{\mu}$ is the characteristicly function of the random variable f . If ρ is the density in the point (x, y) of a mass equal with the unit distributed in plane x, y , then

$$(9) \quad \widehat{\mu}(u, v) = \int_{\mathbb{R}^2} e^{i(ux+vy)} d\mu(x, y) = \int_{\mathbb{R}^2} e^{i(ux+vy)} P(x, y) dx dy$$

Theorem 1 For each measure μ on \mathbb{R}^2 we have:

- i) $\widehat{\mu}(0) = 1$
- ii) $\widehat{\mu}(-t) = \overline{\widehat{\mu}(t)}$
- iii) $\forall a_1, a_2, \dots, a_n \in \mathbb{C}$ and $t_1, t_2, \dots, t_n \in \mathbb{R}^2$ we have

$$\sum_{j,k=1}^n a_j \cdot \overline{a_k} \cdot \widehat{\mu}(t_j - t_k) \geq 0$$

- iv) $\widehat{\mu}$ is a uniformly continuous function.

Proof. i) This follows from (8)

$$\text{ii) } \widehat{\mu}(-t) = \int_{\mathbb{R}^2} e^{i\langle -t, z \rangle} d\mu(z) = \int_{\mathbb{R}^2} \overline{e^{i\langle t, z \rangle}} d\mu(z) = \overline{\widehat{\mu}(t)}$$

$$\begin{aligned} \text{iii) } \sum_{j,k} a_j \overline{a_k} \widehat{\mu}(t_j - t_k) &= \int_{\mathbb{R}^2} \sum_{j,k} a_j \cdot \overline{a_k} e^{i\langle t_j - t_k, z \rangle} d\mu(z) = \\ &= \int_{\mathbb{R}^2} \left| \sum_j a_j \cdot e^{i\langle t, z \rangle} \right|^2 \cdot d\mu(z) \geq 0 \end{aligned}$$

$$\text{iv) } \forall \nu = (h, k) \in \mathbb{R}_+^2, |\widehat{\mu}(u + h, v + k) - \widehat{\mu}(u, v)| \leq \int_{\mathbb{R}^2} |e^{i\langle \nu, z \rangle} - 1| d\mu(z),$$

where $|\nu| = (h^2 + k^2)^{\frac{1}{2}}$.

The integrand is bounded and it tends to zero for $\nu \rightarrow 0$. Then according to Lebesgue's dominated convergence theorem, we have

$$\lim_{|\nu| \rightarrow 0} \sup_{(u,v) \in \mathbb{R}^2} |\widehat{\mu}(u + h, v + k) - \widehat{\mu}(u, v)| = 0$$

and so $\widehat{\mu}$ is uniformly continuous.

Theorem 2 *Let μ be a measure ON \mathbb{R}^2 and*

$$\int |x_j| d\mu(x) < \infty, \quad j = \overline{1, k}, \quad x = (x_1, \dots, x_k), \quad t = (t_1, \dots, t_k).$$

Then $\widehat{\mu}$ is partial derivative with respect to t_j and we have

$$(10) \quad \frac{\partial \widehat{\mu}}{\partial t_j}(t) = i \int x_j \cdot e^{i\langle t, z \rangle} d\mu(x)$$

The partial derivatives $\frac{\partial \widehat{\mu}}{\partial t_j}(t)$ are uniformly continuous.

Proof. Let e_j be the vectors of an orthonormal base. Then we have

$$\frac{\widehat{\mu}(t + h e_j) - \widehat{\mu}(t)}{h} = \int e^{i\langle t, x \rangle} \cdot \frac{e^{i h x_j} - 1}{h} d\mu(x)$$

$$\text{and } \left| e^{i\langle t, x \rangle} \cdot \frac{e^{i h x_j} - 1}{h} \right| \leq |x_j|.$$

For $h \rightarrow 0$ and using the Lebesgue's dominated convergence theorem, we obtain (10). The second part of the theorem follows from Theorem (1).

Observation 1 *It is easy to show that, if μ is a measure on \mathbb{R}^k and*

$$\int |x^n| d\mu(x) < \infty, x^n = x_1^{n_1}, \dots, x_k^{n_k}, n_i \geq 0, |n| = n_1 + \dots + n_k,$$

then the $\frac{\partial^{|n|}}{\partial t_1^{n_1} \dots \partial t_k^{n_k}}(\widehat{\mu}(t))$ exists and

$$(11) \quad \frac{\partial^{|n|}}{\partial t_1^{n_1} \dots \partial t_k^{n_k}} \widehat{\mu}(t) = i^{|n|} \cdot \int x_1^{n_1} \dots x_k^{n_k} \cdot e^{i\langle t, x \rangle} \cdot d\mu(x)$$

where i -imaginary unit.

Theorem 3 *Let μ be a probability measure on \mathbb{R}^2 , so that $\int |x| d\mu(x, y) < \infty, \int |y| d\mu(x, y) < \infty$. Then*

a) If $\int x d\mu(x, y) = 0, \int y d\mu(x, y) = 0$ we have

$$(12) \quad \lim_{n \rightarrow \infty} \left[\widehat{\mu} \left(\frac{u}{n}, \frac{v}{n} \right) \right]^n = 1$$

b) If in addition,

$$\int x^2 d\mu(x, y) = 1, \int y^2 d\mu(x, y) = 1 \text{ and } \int xdy\mu(x, y) = 0$$

we have

$$\lim_{n \rightarrow \infty} \left[\widehat{\mu} \left(\frac{u}{n}, \frac{v}{n} \right) \right]^n = e^{-\frac{1}{2}(u^2+v^2)}$$

Proof. a) From hypothesis and from the theorem (2) follows that $\widehat{\mu}(u, v)$ is differentiable and it's partial derivatives are continuous. Since

$$\frac{\partial \widehat{\mu}(u, v)}{\partial u} = i \int x \cdot e^{i(ux+vy)} d\mu(x, y), \quad \frac{\partial \widehat{\mu}(u, v)}{\partial v} = i \int y e^{i(ux+vy)} d\mu(x, y)$$

follows that

$$\frac{\partial \widehat{\mu}}{\partial u}(0, 0) = i \int x d\mu(x, y) = 0, \quad \frac{\partial \widehat{\mu}}{\partial v}(0, 0) = i \int y d\mu(x, y) = 0$$

Applying the formula Mac-Laurin, we obtain

$$\widehat{\mu}(t) = \widehat{\mu}(u, v) = \widehat{\mu}(0, 0) + u \frac{\partial \widehat{\mu}}{\partial u}(\theta u, \theta v) + v \frac{\partial \widehat{\mu}}{\partial v}(\theta u, \theta v) =$$

$$= 1 + u \cdot \alpha(t) + v \cdot \beta(t)$$

for $|u| \leq 1$, $|v| \leq 1$ where $\alpha(t), \beta(t)$ are continuous functions in $(0,0)$, and $\alpha(0) = 0$, $\beta(0) = 0$, $0 < \theta < 1$. Then,

$$\widehat{\mu} \left(\frac{t}{n} \right) = \widehat{\mu} \left(\frac{u}{n}, \frac{v}{n} \right) = 1 + \frac{u}{n} \alpha \left(\frac{t}{n} \right) + \frac{v}{n} \beta \left(\frac{t}{n} \right) = 1 + \gamma_n(t)$$

where $\gamma_n(t) = \frac{u}{n} \alpha \left(\frac{t}{n} \right) + \frac{v}{n} \beta \left(\frac{t}{n} \right)$ with $\lim_{n \rightarrow \infty} n\gamma_n(t) = 0$, $\forall (u, v) \in \mathbb{R}^2$. Then $\forall (u, v)$, we have

$$\lim_{n \rightarrow \infty} \left[\widehat{\mu} \left(\frac{t}{n} \right) \right]^n = \lim_{n \rightarrow \infty} [1 + \gamma_n(t)]^n = \lim_{n \rightarrow \infty} \left[(1 + \gamma_n(t))^{\frac{1}{\gamma_n(t)}} \right]^{n \cdot \gamma_n(t)} = 1$$

b) From the theorem (2) follows that $\widehat{\mu}$ is twice differentiable with the partial derivatives of the second continuous order and

$$\frac{\partial \widehat{\mu}}{\partial u}(0,0) = 0, \quad \frac{\partial \widehat{\mu}}{\partial v}(0,0) = 0, \quad \frac{\partial^2 \widehat{\mu}}{\partial u^2}(0,0) = -1, \quad \frac{\partial^2 \widehat{\mu}}{\partial v^2}(0,0) = -1, \quad \frac{\partial^2 \widehat{\mu}}{\partial u \partial v}(0,0) = 0.$$

Applying again the formula Mac-Laurin for $|u| \leq 1$, $|v| \leq 1$, $0 < \theta < 1$ we have

$$\begin{aligned} \widehat{\mu}(t) = \widehat{\mu}(u, v) &= \widehat{\mu}(0,0) + u \frac{\partial \widehat{\mu}}{\partial u}(0,0) + v \frac{\partial \widehat{\mu}}{\partial v}(0,0) + \\ &+ \frac{1}{2} \left[u^2 \frac{\partial^2 \widehat{\mu}}{\partial u^2}(\theta u, \theta v) + 2uv \frac{\partial^2 \widehat{\mu}}{\partial u \partial v} + v^2 \cdot \frac{\partial^2 \widehat{\mu}}{\partial v^2}(\theta u, \theta v) \right] \text{ or} \end{aligned}$$

$$\widehat{\mu}(t) = \widehat{\mu}(u, v) = 1 + \frac{1}{2} u^2 \theta_1(t) + uv \theta_2(t) + \frac{1}{2} v^2 \theta_3(t) \text{ where}$$

$$\theta_1(0,0) = -1, \theta_3(0,0) = -1, \theta_2(0,0) = 0. \text{ Then,}$$

$$\begin{aligned} \widehat{\mu} \left(\frac{t}{\sqrt{n}} \right) &= \widehat{\mu} \left(\frac{u}{\sqrt{n}}, \frac{v}{\sqrt{n}} \right) = 1 + \frac{1}{2} \cdot \frac{u^2}{n} \theta_1 \left(\frac{t}{\sqrt{n}} \right) + \frac{uv}{n} \theta_2 \left(\frac{t}{\sqrt{n}} \right) + \\ &+ \frac{1}{2} \cdot \frac{v^2}{n} \theta_3 \left(\frac{t}{\sqrt{n}} \right) = 1 + \sigma_n(t), \text{ where} \end{aligned}$$

$$\sigma_n(t) = \frac{1}{2} \frac{u^2}{n} \theta_1 \left(\frac{t}{\sqrt{n}} \right) + \frac{uv}{n} \theta_2 \left(\frac{t}{\sqrt{n}} \right) + \frac{1}{2} \frac{v^2}{n} \theta_3 \left(\frac{t}{\sqrt{n}} \right)$$

and

$$\lim_{n \rightarrow \infty} n\sigma_n(t) = -\frac{1}{2}(u^2 + v^2), \quad \forall (u, v)$$

Then

$$\lim_{n \rightarrow \infty} \left[\hat{\mu} \left(\frac{t}{\sqrt{n}} \right) \right]^n = e^{-\frac{1}{2}(u^2+v^2)}.$$

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Department of Mathematics
University "Lucian Blaga" of Sibiu,
Str. Dr. I. Ratiu Nr.7
2400 Sibiu, Romania